



# Probability, Third Edition

By David J Carr & Michael A Gauger

Published by BPP Professional Education

## Solutions to practice questions – Chapter 7

### Solution 7.1

- (i)  $\Pr(X > 3.1) = 1 - \Phi(3.1) = 1 - 0.9990 = 0.0010$
- (ii)  $\Pr(X < -1.4) = \Phi(-1.4) = 1 - \Phi(1.4) = 1 - 0.9192 = 0.0808$
- (iii)  $\Pr(0.4 < X < 2.2) = \Phi(2.2) - \Phi(0.4) = 0.9861 - 0.6554 = 0.3307$
- (iv)  $\Pr(-1.7 < X < -0.2) = \Phi(-0.2) - \Phi(-1.7) = (1 - \Phi(0.2)) - (1 - \Phi(1.7))$   
 $= \Phi(1.7) - \Phi(0.2) = 0.9554 - 0.5793 = 0.3761$

### Solution 7.2

The moment generating function of  $X$  is:

$$M_X(t) = \exp\left(\mu t + \frac{\sigma^2 t^2}{2}\right)$$

Hence the cumulant generating function of  $X$  is:

$$R_X(t) = \ln(M_X(t)) = \mu t + \frac{\sigma^2 t^2}{2}$$

- (i) The expected value of  $X$  is:

$$E[X] = R'_X(0) = \mu + 0 \times \sigma^2 = \mu$$

- (ii) The variance of  $X$  is:

$$\text{var}(X) = R''_X(0) = \sigma^2$$

**Solution 7.3**

The moment generating function is:

$$\exp(8(t^2 - 1.5t)) = \exp\left(-12t + \frac{16t^2}{2}\right)$$

Comparing this to the MGF of a normal distribution:

$$M_X(t) = \exp\left(\mu t + \frac{\sigma^2 t^2}{2}\right)$$

We can see that  $X \sim N(-12, 16)$ .

**Solution 7.4**

By Theorem 7.1,  $\left(\frac{X-15}{\sqrt{100}}\right) \sim N(0,1)$ . Hence:

$$(i) \quad \Pr(X > 18) = 1 - \Pr(X \leq 18) = 1 - \Pr\left(\frac{X-15}{10} \leq \frac{18-15}{10}\right) = 1 - \Phi(0.3) = 1 - 0.6179 = 0.3821$$

$$(ii) \quad \Pr(3 < X < 27) = \Pr\left(\frac{3-15}{10} < \frac{X-15}{10} < \frac{27-15}{10}\right) = \Phi(1.2) - \Phi(-1.2) \\ = \Phi(1.2) - (1 - \Phi(1.2)) = 2 \times \Phi(1.2) - 1 = 2 \times 0.8849 - 1 = 0.7698$$

$$(iii) \quad \Pr(-4 < X < 4) = \Pr\left(\frac{-4-15}{10} < \frac{X-15}{10} < \frac{4-15}{10}\right) = \Phi(-1.1) - \Phi(-1.9) \\ = (1 - \Phi(1.1)) - (1 - \Phi(1.9)) = \Phi(1.9) - \Phi(1.1) = 0.9713 - 0.8643 = 0.1070$$

**Solution 7.5**

The probability that a sample value ( $X$ ) is less than 9 is:

$$\Pr(X < 9) = \Pr\left(\frac{X-10}{\sqrt{2}} < \frac{9-10}{\sqrt{2}}\right) = \Phi(-0.7071) = 1 - \Phi(0.7071) = 1 - 0.7601 = 0.2399$$

This last value is found using linear interpolation:

$$\Phi(0.7071) \approx \Phi(0.7) + \frac{0.7071 - 0.7}{0.8 - 0.7} \times [\Phi(0.8) - \Phi(0.7)] = 0.7601$$

Hence, the required probability is:

$$({}_6C_2)(0.2399)^2 (0.7601)^4 = 0.2882$$

**Solution 7.6**

Let  $X_1$  and  $X_2$  be the sizes of the two claims, and let  $Y = X_1 - X_2$  be the difference between these random variables. Then by Theorem 7.4 (with  $a_1 = 1$  and  $a_2 = -1$ ), we have:

$$Y \sim N\left((1,800 - 1,800), (400^2 + 400^2)\right) = N(0, 320,000)$$

The probability that the claims differ by more than \$500 is:

$$\Pr(Y < -500) + \Pr(Y > 500) = 2 \times \Pr(Y > 500) \quad \text{by symmetry}$$

We have:

$$\begin{aligned} \Pr(Y > 500) &= 1 - \Pr(Y \leq 500) = 1 - \Pr\left(\frac{Y - 0}{\sqrt{320,000}} \leq \frac{500 - 0}{\sqrt{320,000}}\right) \\ &= 1 - \Phi(0.8839) = 1 - 0.8114 = 0.1886 \end{aligned}$$

Hence the required probability is:

$$\Pr(Y < -500) + \Pr(Y > 500) = 2 \times \Pr(Y > 500) = 2 \times 0.1886 = 0.3772$$

**Solution 7.7**

Let  $X_i$  be the lifetime of the  $i$ -th bulb. A succession of  $n$  bulbs will produce light for a random time equal to:

$$S = X_1 + \dots + X_n$$

Since  $X_i \sim N(3, 1)$  and the lifetimes are independent, we know that  $S$  is normally distributed and:

$$E[S] = nE[X] = 3n$$

$$\text{var}(S) = n \text{var}(X) = n$$

We want to find the smallest  $n$  such that  $\Pr(S > 40) \geq 0.9772$ .

The first step is to convert this probability statement to one about the standard normal distribution:

$$\Pr(S > 40) = \Pr\left(\frac{S - 3n}{\sqrt{n}} > \frac{40 - 3n}{\sqrt{n}}\right) = 1 - \Phi\left(\frac{40 - 3n}{\sqrt{n}}\right)$$

From the normal distribution table, note that:

$$\Phi(2) = 0.9772$$

$$\Rightarrow 1 - \Phi(-2) = 0.9772 \quad (\text{by symmetry})$$

Hence:

$$\frac{40 - 3n}{\sqrt{n}} = -2 \quad \Rightarrow 3n - 2\sqrt{n} - 40 = 0$$

We can view this equation as a quadratic in the variable  $\sqrt{n}$ .

By the quadratic formula:

$$\begin{aligned}\sqrt{n} &= \frac{-(-2) \pm \sqrt{4 - 4 \times 3 \times (-40)}}{2 \times 3} = \frac{2 \pm 22}{6} \\ \Rightarrow \sqrt{n} &= \frac{24}{6} = 4 \quad (\text{positive root}) \\ \Rightarrow n &= 16\end{aligned}$$

**Note:** If, for example, the solution of the quadratic equation were  $n = 16.13$ , then the answer would be  $n = 17$  and the probability would be slightly higher than 0.9772.

### Solution 7.8

Let the less accurate measurement be  $X$  and let the more accurate measurement be  $Y$ . Then:

$$\begin{aligned}X &\sim N\left(h, (0.0056h)^2\right) \\ Y &\sim N\left(h, (0.0044h)^2\right)\end{aligned}$$

Let  $Z$  be the average value of  $X$  and  $Y$ . Then  $Z$  is normally distributed with:

$$\begin{aligned}\mu &= E[Z] = E\left[\frac{X+Y}{2}\right] = \frac{1}{2}(E[X] + E[Y]) = \frac{1}{2}(h+h) = h \\ \sigma^2 &= \text{var}(Z) = \text{var}\left(\frac{X+Y}{2}\right) = \left(\frac{1}{2}\right)^2 \text{var}(X+Y) = \frac{1}{4}(\text{var}(X) + \text{var}(Y)) = 0.00001268h^2 \\ \Rightarrow \sigma &= 0.003561h\end{aligned}$$

To calculate the probability that the average value is within  $0.005h$  of the actual height  $h$ :

$$\begin{aligned}\Pr(0.995h < Z < 1.005h) &= \Pr\left(\frac{0.995h-h}{0.003561h} < \frac{Z-h}{0.003561h} < \frac{1.005h-h}{0.003561h}\right) \\ &= \Pr\left(-1.4041 < \frac{Z-h}{0.003561h} < 1.4041\right) \\ &= \Phi(1.4041) - \Phi(-1.4041) \\ &= \Phi(1.4041) - (1 - \Phi(1.4041)) \\ &= 2\Phi(1.4041) - 1 = 2 \times 0.9198 - 1 = 0.8396\end{aligned}$$

**Solution 7.9**

Let  $X_i$  be the amount of the  $i$ -th claim. The total claim amount for the 200 claims is:

$$S = X_1 + X_2 + \dots + X_{200}$$

Since  $X_i \sim N(237, 202^2)$  and the claims are independent, then  $S$  is normally distributed with:

$$E[S] = 200 \times 237 = 47,400$$

$$\text{var}(S) = 200 \times 202^2 = 8,160,800$$

The required probability is:

$$\begin{aligned} \Pr(S > 50,000) &= 1 - \Pr(S \leq 50,000) = 1 - \Pr\left(\frac{S - 47,400}{\sqrt{8,160,800}} \leq \frac{50,000 - 47,400}{\sqrt{8,160,800}}\right) \\ &= 1 - \Phi(0.9101) = 1 - 0.8185 = 0.1815 \end{aligned}$$

**Solution 7.10**

Let  $X_i$  be the amount of the  $i$ -th claim. The average amount of the 80 claims is:

$$\bar{X} = \frac{1}{80}(X_1 + X_2 + \dots + X_{80})$$

Since  $X_i \sim N(574, 186^2)$  and the claims are independent, then  $\bar{X}$  is normally distributed with:

$$E[\bar{X}] = 574 \quad \text{var}(\bar{X}) = \frac{1}{80} \times 186^2 = 432.45$$

The required probability is:

$$\Pr(\bar{X} < 565) = \Pr\left(\frac{\bar{X} - 574}{\sqrt{432.45}} \leq \frac{565 - 574}{\sqrt{432.45}}\right) = \Phi(-0.4328) = 1 - \Phi(0.4328) = 1 - 0.6672 = 0.3328$$

**Solution 7.11**

You should recognize that the given density is exponential with  $\theta = 1,000$ . Therefore, the mean and variance of the claim amount per policy,  $X$ , are:

$$E[X] = \theta = 1,000 \quad \text{var}(X) = \theta^2 = 1,000,000$$

The total claims for the group of 100 policies can be written as:

$$S = X_1 + \dots + X_{100}$$

where the various  $X_i$  are identically distributed like  $X$ .

The phrase in the question “the *approximate* probability” is a clue to use the Central Limit Theorem, even though the question does not state that the claim amounts for different policies are independent. If we assume independence, then:

$$E[S] = 100E[X] = 100,000$$

$$\text{var}(S) = 100\text{var}(X) = 100,000,000 \quad \Rightarrow \quad \sigma_S = \sqrt{\text{var}(S)} = 10,000$$

By the Central Limit Theorem,  $S$  is approximately normal in distribution and:

$$\frac{S - E[S]}{\sigma_S} = \frac{S - 100,000}{10,000} \sim N(0, 1)$$

The premium for a single policy is:

$$100 + E[X] = 100 + 1,000 = 1,100$$

The premium for the group of 100 policies is thus 110,000. So, the probability that claims exceed premium is given by:

$$\begin{aligned} \Pr(S > 110,000) &= \Pr\left(\frac{S - 100,000}{10,000} > \frac{110,000 - 100,000}{10,000}\right) \\ &= \Pr\left(\frac{S - 100,000}{10,000} > 1\right) \approx 1 - \Phi(1) = 1 - 0.8413 = 0.1587 \end{aligned}$$

### Solution 7.12

We can make use of results in Solution 7.11.

Let  $G$  be the group premium. The value of  $G$  is set such that:

$$\begin{aligned} \Pr(S > G) &= 0.1 \\ \Rightarrow \Pr(S \leq G) &= 0.9 \\ \Rightarrow \Pr\left(\frac{S - 100,000}{10,000} \leq \frac{G - 100,000}{10,000}\right) &= 0.9 \\ \Rightarrow \Phi\left(\frac{G - 100,000}{10,000}\right) &= 0.9 \\ \Rightarrow \frac{G - 100,000}{10,000} &= 1.282 \end{aligned}$$

Solving for  $G$  allows us to calculate the premium per policy:

$$\begin{aligned} \frac{G - 100,000}{10,000} = 1.282 &\Rightarrow G = 112,820 \\ \Rightarrow \text{Premium per policy} &= \frac{G}{100} = 1,128.20 \end{aligned}$$

### Solution 7.13

Let  $X_i$  be the difference between the true age and the rounded age of the  $i$ -th individual.

Since  $X_i$  is distributed uniformly on the interval  $[-2.5, 2.5]$ , it follows that:

$$\begin{aligned} E[X_i] &= \frac{-2.5 + 2.5}{2} = 0 \\ \text{var}(X_i) &= \frac{(2.5 - (-2.5))^2}{12} = \frac{25}{12} = 2.08333 \end{aligned}$$

Assuming that  $X_1, X_2, \dots, X_{48}$  are independent, then using the Central Limit Theorem:

$$\bar{X} = \frac{1}{48} \sum_{i=1}^{48} X_i \text{ is approximately distributed } N\left(0, \frac{1}{48} \times 2.08333\right) = N(0, 0.0434)$$

Finally, we have:

$$\begin{aligned} \Pr(-0.25 < \bar{X} < 0.25) &= \Pr\left(\frac{-0.25-0}{\sqrt{0.0434}} < \frac{\bar{X}-0}{\sqrt{0.0434}} < \frac{0.25-0}{\sqrt{0.0434}}\right) = \Pr\left(-1.2 < \frac{\bar{X}-0}{\sqrt{0.0434}} < 1.2\right) \\ &\approx \Phi(1.20) - \Phi(-1.20) = \Phi(1.20) - (1 - \Phi(1.20)) \\ &= 2\Phi(1.20) - 1 = 2 \times 0.8849 - 1 = 0.7698 \end{aligned}$$

### Solution 7.14

The number of sales,  $N$ , follows a binomial distribution with parameters  $n = 200$  and  $p = 0.7$ . Hence:

$$E[N] = np = 140 \quad \text{var}(N) = npq = 42$$

We can approximate the distribution of  $N$  with the random variable  $X \sim N(140, 42)$ .

Hence the required probability is:

$$\begin{aligned} \Pr(124 < N \leq 153) &\approx \Pr(124.5 < X < 153.5) = \Pr\left(\frac{124.5-140}{\sqrt{42}} < \frac{X-140}{\sqrt{42}} < \frac{153.5-140}{\sqrt{42}}\right) \\ &= \Phi(2.0831) - \Phi(-2.3917) = \Phi(2.0831) - (1 - \Phi(2.3917)) \\ &= 0.9813 - (1 - 0.9916) = 0.9729 \end{aligned}$$

### Solution 7.15

By the additive property of the Poisson distribution, the total number of claims,  $N$ , follows a Poisson distribution with mean  $0.3 \times 1,000 = 300$ . Hence:

$$E[N] = 300 \quad \text{var}(N) = 300$$

We can approximate the distribution of  $N$  with the random variable  $X \sim N(300, 300)$ .

Hence the required probability is:

$$\begin{aligned} \Pr(N \geq 340) &\approx \Pr(X > 339.5) = 1 - \Pr(X \leq 339.5) = 1 - \Pr\left(\frac{X-300}{\sqrt{300}} \leq \frac{339.5-300}{\sqrt{300}}\right) \\ &= 1 - \Phi(2.2805) = 1 - 0.9887 = 0.0113 \end{aligned}$$

### Solution 7.16

$Y$  follows a lognormal distribution with parameters  $\mu = 0.5$  and  $\sigma^2 = 0.16$ .

The expected value of  $Y$  is:

$$E[Y] = \exp\left(\mu + \frac{1}{2}\sigma^2\right) = \exp\left(0.5 + \frac{1}{2} \times 0.16\right) = \exp(0.58) = 1.786$$

**Solution 7.17**

The variance of is:

$$\begin{aligned}\text{var}(X) &= \exp(2\mu + 2\sigma^2) - \exp(2\mu + \sigma^2) \\ &= \exp(2 \times 2.5 + 2 \times 1.6^2) - \exp(2 \times 2.5 + 1.6^2) \\ &= \exp(10.12) - \exp(7.56) = 22,914.93\end{aligned}$$

**Solution 7.18**

By Theorem 7.8,  $e^X$  follows a lognormal distribution with parameters  $\mu = 0.06$  and  $\sigma = 0.015$ .

The expected value of the savings account in one year's time is:

$$E[Y] = 10,000 \exp\left(\mu + \frac{1}{2}\sigma^2\right) = 10,000 \exp\left(0.06 + \frac{0.015^2}{2}\right) = 10,619.56$$

**Solution 7.19**

By the additive property for the chi-square distribution (covered on page 176),  $Y \sim \chi^2$  with 10 degrees of freedom.

Hence the pdf of  $Y$  is:

$$f(y) = \frac{y^{(r/2)-1} e^{-y/2}}{2^{r/2} \Gamma(r/2)} = \frac{y^4 e^{-y/2}}{2^5 \Gamma(5)} = \frac{1}{768} y^4 e^{-y/2}$$

**Solution 7.20**

Let  $X_1$  be the horizontal error, and let  $X_2$  be the vertical error. The distance from the center of the target is equal to:

$$\sqrt{X_1^2 + X_2^2}$$

Since  $X_i \sim N(0, 1.25^2)$ , then by Theorem 7.10 we have:

$$\left(\frac{X_i - \mu}{\sigma}\right)^2 = \left(\frac{X_i}{1.25}\right)^2 \sim \chi^2 \quad \text{with 1 degree of freedom}$$

Let  $Y = \left(\frac{X_1}{1.25}\right)^2 + \left(\frac{X_2}{1.25}\right)^2$ . Then by the additive property, we have:

$$Y \sim \chi^2 \quad \text{with 2 degrees of freedom}$$

So, the required probability is:

$$\Pr\left(\sqrt{X_1^2 + X_2^2} < 1\right) = \Pr\left(X_1^2 + X_2^2 < 1\right) = \Pr\left(\left(\frac{X_1}{1.25}\right)^2 + \left(\frac{X_2}{1.25}\right)^2 < \frac{1}{1.25^2}\right) = \Pr(Y < 0.64)$$

Using the property that a chi-square distribution with parameter  $r=2$  is equivalent to the exponential distribution with parameter  $\theta=2$ .

Hence:

$$\Pr\left(\sqrt{X_1^2 + X_2^2} < 1\right) = \Pr(Y < 0.64) = 1 - e^{-0.64/2} = 0.2739$$