



Probability, Second Edition

By David J Carr & Michael A Gauger Published by BPP Professional Education

Solutions to practice questions - Chapter 9

Solution 9.1

Let *X* be the number of demands made each day, and let *Y* be the number of demands handled.

Then we have:

$$Y = \min\{X,3\}$$

and:

$$Pr(Y=0) = Pr(X=0) = e^{-2} = 0.13534$$

$$Pr(Y=1) = Pr(X=1) = 2e^{-2} = 0.27067$$

$$Pr(Y=2) = Pr(X=2) = \frac{2^2 e^{-2}}{2!} = 0.27067$$

$$Pr(Y=3) = Pr(X \ge 3) = 1 - 0.13534 - 0.27067 - 0.27067 = 0.32332$$

Hence, the expected number of demands handled is:

$$E[Y] = 0 \times 0.13534 + 1 \times 0.27067 + 2 \times 0.27067 + 3 \times 0.32332$$

= 1.782

Solution 9.2

We can calculate E[Y] and $E[Y^2]$ in terms of E[X] and $E[X^2]$:

$$E[Y] = \sum_{y=0}^{\infty} y \Pr(Y = y) = \sum_{y=1}^{\infty} y \Pr(Y = y) = (1 - \alpha) \sum_{y=1}^{\infty} y \Pr(X = y) = (1 - \alpha) E[X] = (1 - \alpha) \mu$$

$$E[Y^{2}] = \sum_{y=0}^{\infty} y^{2} \Pr(Y = y) = \sum_{y=1}^{\infty} y^{2} \Pr(Y = y) = (1 - \alpha) \sum_{y=1}^{\infty} y^{2} \Pr(X = y) = (1 - \alpha) E[X^{2}] = (1 - \alpha) (\mu^{2} + \mu)$$

Hence, the variance of Y is:

$$var(Y) = E[Y^{2}] - (E[Y])^{2}$$

$$= (1 - \alpha)(\mu^{2} + \mu) - ((1 - \alpha)\mu)^{2}$$

$$= (1 - \alpha)\mu(1 + \alpha\mu)$$

Note that since the pdf is defined for x > 0, the transformation $Y = X^2$ is 1-1, with:

$$X = +\sqrt{Y}$$

So, using the method of transformations:

$$f_Y(y) = f_X(\sqrt{y}) \left| \frac{d(\sqrt{y})}{dy} \right| = e^{-\sqrt{y}} \times \frac{1}{2\sqrt{y}}$$
 for $y > 0$

Solution 9.4

The pdf of X is:

$$f_X(x) = \frac{x^{\alpha - 1}e^{-x/\theta}}{\theta^{\alpha}\Gamma(\alpha)}$$
 for $x > 0$

Now, if Y = g(X) = cX, we have a 1-1 differentiable transformation.

The inverse transformation is:

$$X = Y / c$$

So, using the method of transformations:

$$f_Y(y) = f_X\left(\frac{y}{c}\right) \left| \frac{d(y/c)}{dy} \right| = \frac{(y/c)^{\alpha - 1} e^{-(y/c)/\theta}}{\theta^{\alpha} \Gamma(\alpha)} \times \frac{1}{c}$$
$$= \frac{y^{\alpha - 1} e^{-y/(c\theta)}}{(c\theta)^{\alpha} \Gamma(\alpha)} \quad \text{for } y > 0$$

Hence *Y* follows a gamma distribution with parameters α and $c\theta$.

Solution 9.5

Let's use the method of distribution functions.

We have:

$$F_Y(y) = \Pr(Y \le y) = \Pr\left(\frac{1}{X} - 1 \le y\right) = \Pr\left(X \ge \frac{1}{1 + y}\right) = 1 - \Pr\left(X < \frac{1}{1 + y}\right) = 1 - \frac{1}{1 + y}$$

Hence, the pdf of Y is:

$$f_Y(y) = F'_Y(y) = \frac{1}{(1+y)^2}$$
 for $y > 0$

The density of T is:

$$f(t) = \frac{1}{3}e^{-t/3}$$
 for $t > 0$

The discovery time is a piecewise function of T:

$$X = g(T) = \max(T, 2) = \begin{cases} 2 & T \le 2 \\ T & T > 2 \end{cases}$$

Hence:

$$E[X] = \int_0^\infty g(t) f(t) dt = 2 \int_0^2 \frac{e^{-t/3}}{3} dt + \int_2^\infty \frac{t e^{-t/3}}{3} dt$$

$$= 2 \times F(2) + \left(-t e^{-t/3} - 3 e^{-t/3} \right) \Big|_2^\infty \qquad \text{(integrating by parts)}$$

$$= 2 \left(1 - e^{-2/3} \right) + \left(-0 - 0 + 2 e^{-2/3} + 3 e^{-2/3} \right)$$

$$= 2 + 3 e^{-2/3}$$

Solution 9.7

We are given that *T* is distributed uniformly on the interval [8,12], so the probability density function is:

$$f_T(t) = \frac{1}{4}$$
 for $8 \le t \le 12$

The random rate of service in customers per minute is:

$$R = 10 / T$$

Since this is a 1-1, differentiable transformation of T with inverse function T = 10/R, we can calculate the probability density function $f_R(r)$ using the method of transformations:

$$f_R(r) = f_T(g^{-1}(r)) \left| \frac{dg^{-1}(r)}{dr} \right|$$
$$= f_T(\frac{10}{r}) \times \left| \frac{d(10/r)}{dr} \right|$$
$$= \frac{1}{4} \times \left| \frac{-10}{r^2} \right|$$
$$= \frac{5}{2r^2}$$

Let *Y* be the length of time lived in the next 10 years.

Then *Y* is defined as follows:

$$Y = \min(X, 10) = \begin{cases} X & 0 \le X \le 10 \\ 10 & X > 10 \end{cases}$$

The expected value of Y is therefore:

$$E[Y] = \int g(x) f(x) dx$$

$$= \int_0^{10} x \frac{1}{75} dx + \int_{10}^{75} 10 \frac{1}{75} dx$$

$$= \frac{x^2}{150} \Big|_0^{10} + \frac{10x}{75} \Big|_{10}^{75}$$

$$= 9.333$$

Note: This question uses the theory related to policy limits, since we apply a maximum of 10 to the future lifetime of the 25-year-old.

Solution 9.9

The benefit amount paid by the insurer is:

$$Y = \max\{0, (X - 250)\} = \begin{cases} 0 & 0 \le X \le 250 \\ X - 250 & X > 250 \end{cases}$$

So the expected benefit amount is:

$$E[Y] = \int_0^{250} 0 \times f_X(x) dx + \int_{250}^{\infty} (x - 250) f_X(x) dx$$
$$= 0 + \int_{250}^{\infty} (x - 250) \frac{1}{1,000} e^{-x/1,000} dx$$

We could use integration by parts to finish the calculation. However, it is quite a bit quicker to employ the substitution z = x - 250:

$$E[Y] = \int_0^\infty z \frac{1}{1,000} e^{-(z+250)/1,000} dz$$
$$= e^{-250/1,000} \int_0^\infty z \frac{1}{1,000} e^{-z/1,000} dz$$

Finally, note that the last integral is the expected value of an exponential distribution with mean 1,000. So without performing any further integration, we can see that the final result is:

$$E[Y] = e^{-0.25} \times 1,000 = 778.80$$

Let *X* be the manufacturer's annual losses, and let *Y* be the part of the annual losses not paid by the insurance company. Then:

$$Y = \min(X, 2) = \begin{cases} X & 0.6 < X \le 2 \\ 2 & X > 2 \end{cases}$$

We can calculate the expected value of Y as:

$$E[Y] = \int_{0.6}^{2} x f(x) dx + \int_{2}^{\infty} 2 f(x) dx = \int_{0.6}^{2} x f(x) dx + 2 \times Pr(X > 2)$$

The first component requires simple integration:

$$\int_{0.6}^{2} x f(x) dx = 2.5(0.6)^{2.5} \int_{0.6}^{2} x^{-2.5} dx = -\frac{2.5(0.6)^{2.5} x^{-1.5}}{1.5} \bigg|_{0.6}^{2} = 0.8357$$

The second component can be calculated most efficiently by observing that X follows a one-parameter Pareto distribution with $\alpha = 0.25$, $\theta = 0.6$. Then:

$$2 \times \Pr(X > 2) = 2 \times \left[1 - F(2)\right] = 2 \times \left[1 - \left(1 - \left(\frac{\theta}{x}\right)^{\alpha}\right)\right] = 2 \times \left(\frac{0.6}{2}\right)^{2.5} = 0.0986$$

Hence:

$$E[Y] = 0.8357 + 0.0986 = 0.9343$$

Solution 9.11

Since Y = 1.05 X is a 1-1 differentiable transformation, we can use the method of transformations.

First, we identify the inverse function of the transformation. This is simply:

$$x = y / 1.05$$

Hence, the density function of Y, the loss in 2003, is given by:

$$f_Y(y) = f_X(g^{-1}(y)) \left| \frac{dg^{-1}(y)}{dy} \right| = f_X(y/1.05) \left| \frac{d(y/1.05)}{dy} \right|$$
$$= \frac{2 \times 1,000^2}{\left((y/1.05) + 1,000 \right)^3} \times \frac{1}{1.05}$$
$$= \frac{2 \times 1,050^2}{\left(y + 1,050 \right)^3} \quad \text{for } y > 0$$

Let $Y = X_1 + X_2$ where X_1 and X_2 are the lifetimes of the primary and backup systems.

The joint density of X_1 and X_2 is:

$$f(x_1, x_2) = f(x_1) f(x_2) = 0.05 e^{-0.05x_1} \times 0.10 e^{-0.10x_2}$$
 where $x_1, x_2 > 0$

Step 1: The region in the $X_1 \times X_2$ plane defined by the inequality $X_1 + X_2 \le y$ is a triangular region in the first quadrant below the line $X_1 + X_2 = y$.

Step 2: The cdf of Y is calculated as:

$$\begin{split} F_Y\left(y\right) &= \Pr\left(X_1 + X_2 \le y\right) = \int_{x_2=0}^y \int_{x_1=0}^{y-x_2} 0.05 \, e^{-0.05x_1} \, 0.10 \, e^{-0.10x_2} \, dx_1 \, dx_2 \\ &= \int_{x_2=0}^y 0.10 \, e^{-0.10x_2} \left(1 - e^{-0.05(y-x_2)}\right) dx_2 \\ &= \int_0^y 0.10 e^{-0.10x_2} \, dx_2 - 2e^{-0.05y} \int_0^y 0.05 e^{-0.05x_2} \, dx_2 \\ &= \left(1 - e^{-0.10y}\right) - 2e^{-0.05y} \left(1 - e^{-0.05y}\right) \\ &= 1 + e^{-0.10y} - 2e^{-0.05y} \end{split}$$

Step 3: Differentiate this to obtain the density function of *Y* :

$$f_Y(y) = F'_Y(y) = 0.1(e^{-0.05y} - e^{-0.10y})$$

Solution 9.13

For a Pareto distribution with parameters $\alpha = 2$ and $\theta = 50$, we have:

$$f(x) = \frac{\alpha \theta^{\alpha}}{\left(x+\theta\right)^{\alpha+1}} = \frac{5,000}{\left(x+50\right)^3}$$
$$F(x) = 1 - \left(\frac{\theta}{x+\theta}\right)^{\alpha} = 1 - \left(\frac{50}{x+50}\right)^2$$

Hence, the pdf of the 3rd largest claim (which is also the 23rd smallest claim) is:

$$f_{Y_{23}}(y) = \frac{25!}{22!2!} (F_X(y))^{22} (1 - F_X(y))^2 f_X(y)$$

$$= 6,900 \times \left[1 - \left(\frac{50}{y + 50} \right)^2 \right]^{22} \times \left[\left(\frac{50}{y + 50} \right)^2 \right]^2 \times \frac{5,000}{(y + 50)^3}$$

$$= \frac{2.15625 \times 10^{14}}{(y + 50)^7} \times \left[1 - \left(\frac{50}{y + 50} \right)^2 \right]^{22}$$

For an exponential distribution with mean 10, we have:

$$f(x) = 0.1e^{-0.1x}$$
$$F(x) = 1 - e^{-0.1x}$$

Hence, the pdf of the 4th largest claim (which is also the 2nd smallest claim) is:

$$f_{Y_2}(y) = \frac{5!}{1!3!} (F_X(y))^1 (1 - F_X(y))^3 f_X(y)$$
$$= 20 \times (1 - e^{-0.1y}) \times (e^{-0.1y})^3 \times 0.1e^{-0.1y}$$
$$= 2(e^{-0.4y} - e^{-0.5y})$$

Hence the expected value is:

$$\int_0^\infty y f_{Y_2}(y) dy = \int_0^\infty 2y \left(e^{-0.4y} - e^{-0.5y}\right) dy$$
$$= 5 \int_0^\infty 0.4y e^{-0.4y} dy - 4 \int_0^\infty 0.5y e^{-0.5y} dy$$

Notice that these two integrals are the expected values of exponential distributions with means 2.5 and 2, hence:

$$E[Y_2] = 5 \times 2.5 - 4 \times 2 = 4.5$$

Solution 9.15

The first step here is to calculate the density function of:

$$Y = \max\{X_1, X_2, X_3\}$$

We can use the method of distribution functions to calculate the distribution function of *Y* as follows:

$$F_Y(y) = \Pr(Y \le y) = \Pr(all \ X_i \le y) = \left(F_X(y)\right)^3$$
$$= \left(\int_1^y \frac{3}{x^4} dx\right)^3 = \left(1 - \frac{1}{y^3}\right)^3 \quad y > 1$$

We differentiate the cdf to find the pdf:

$$f_Y(y) = F_Y'(y) = 3\left(1 - \frac{1}{y^3}\right)^2 \frac{3}{y^4} = 9\left(\frac{1}{y^4} - \frac{2}{y^7} + \frac{1}{y^{10}}\right)$$
 $y > 1$

and then compute E[Y] as:

$$E[Y] = \int_{1}^{\infty} y \, f_{Y}(y) \, dy = 9 \int_{1}^{\infty} \frac{1}{y^{3}} - \frac{2}{y^{6}} + \frac{1}{y^{9}} \, dy$$
$$= 9 \left(-\frac{1}{2y^{2}} + \frac{2}{5y^{5}} - \frac{1}{8y^{8}} \right) \Big|_{1}^{\infty} = 9 \left(\frac{1}{2} - \frac{2}{5} + \frac{1}{8} \right) = 2.025 \text{ (thousand)}$$