



# **Probability**, Third Edition

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# Solutions to practice questions – Chapter 6

#### Solution 6.1

Let the losses due to storm, fire, and theft be denoted  $X_1$ ,  $X_2$  and  $X_3$  respectively.

$$Pr(\max\{X_1, X_2, X_3\} > 2.5) = 1 - Pr(\max\{X_1, X_2, X_3\} \le 2.5)$$
$$= 1 - Pr(X_1 \le 2.5) \times Pr(X_2 \le 2.5) \times Pr(X_3 \le 2.5)$$

The distribution function for a uniform distribution with range [a,b] is:

$$F(x) = \frac{x-a}{b-a}$$

So, we can calculate the probability as:

$$\Pr\left(\max\{X_1, X_2, X_3\} > 2.5\right) = 1 - \Pr(X_1 \le 2.5) \times \Pr(X_2 \le 2.5) \times \Pr(X_3 \le 2.5)$$
$$= 1 - (1) \left(\frac{2.5}{3}\right) \left(\frac{2.5}{4}\right) = 1 - 0.521 = 0.479$$

### Solution 6.2

We have:

$$E[X] = \frac{a+b}{2} = 1,879$$
  
var(X) =  $\frac{(b-a)^2}{12} = 507 \implies b-a = \sqrt{12 \times 507} = 78$ 

Hence:

$$a+b = 2 \times 1,879 = 3,758$$
  

$$\Rightarrow a + (78 + a) = 3,758$$
  

$$\Rightarrow a = 1,840$$
  

$$\Rightarrow b = 1,918$$

⇒

From 2010, this question is no longer covered by the Exam P syllabus.

## Solution 6.4

We need:

$$E[X] = \int_{0}^{\infty} x \, 3 \times 2,000^{3} \, (x+2,000)^{-4} \, dx$$

Integrating by parts:

$$E[X] = -\frac{x2,000^3}{(x+2,000)^3} \bigg|_0^{\infty} + \int_0^{\infty} 2,000^3 (x+2,000)^{-3} dx$$
$$= 0 - \frac{2,000^3}{2(x+2,000)^2} \bigg|_0^{\infty}$$
$$= 1,000$$

# Solution 6.5

We want to calculate the following:

$$\Pr(X < 2 | X > 1.5) = \frac{\Pr(X < 2 \cap X > 1.5)}{\Pr(X > 1.5)} = \frac{\Pr(1.5 < X < 2)}{\Pr(X > 1.5)}$$

Now:

$$\Pr(X > 1.5) = \int_{1.5}^{\infty} 3x^{-4} dx = -\frac{1}{x^3} \Big|_{1.5}^{\infty} = \frac{1}{1.5^3} = 0.296296$$
$$\Pr(1.5 < X < 2) = \int_{1.5}^{2} 3x^{-4} dx = -\frac{1}{x^3} \Big|_{1.5}^{2} = \frac{1}{1.5^3} - \frac{1}{2^3} = 0.171296$$

So:

$$\Pr(X < 2 | X > 1.5) = \frac{0.171296}{0.296296} = 0.578125 \quad \text{(exactly)}$$

First we need the value of c. This is found from:

$$\int_{3}^{\infty} c x^{-4} dx = 1$$
$$\Rightarrow \frac{1}{c} = -\frac{1}{3x^{3}} \Big|_{3}^{\infty} = \frac{1}{81}$$
$$\Rightarrow c = 81$$

The expected lifetime is then:

$$E[X] = \int_{3}^{\infty} x \, 81 \, x^{-4} \, dx = -\frac{81}{2 \, x^{2}} \bigg|_{3}^{\infty} = 4.5$$

#### Solution 6.7

For the exponential distribution with mean 50, we have:

$$\Pr(X > x) = 1 - F(x) = e^{-x/50}$$

Hence:

$$\Pr(X > 200 | X > 50) = \frac{\Pr(X > 200)}{\Pr(X > 50)} = \frac{e^{-200/50}}{e^{-50/50}} = \frac{e^{-4}}{e^{-1}} = e^{-3} = 0.0498$$

Alternatively, using the memoryless property we have:

$$\Pr(X > 200 | X > 50) = \Pr(X > 150) = e^{-150/50} = e^{-3} = 0.0498$$

#### Solution 6.8

Let *X* be the random time (in days) until the high-risk driver has an accident. Since *X* follows an exponential distribution, the cdf is:

 $F(x) = 1 - e^{-x/\theta}$ 

We are given that

 $0.30 = \Pr(X < 50) = 1 - e^{-50/\theta}$  $\Rightarrow e^{-50/\theta} = 0.70$ 

Hence:

$$\Pr(X < 80) = 1 - e^{-80/\theta} = 1 - \left(e^{-50/\theta}\right)^{8/5} = 1 - 0.70^{8/5} = 0.4349$$

**Note:** It is unnecessary to calculate the precise value of  $\theta$ . If you did solve for  $\theta$ , you should find that  $\theta = 140.184$ .

Let *T* denote the random time required to repair the machine. It is assumed to be exponentially distributed with mean  $\theta_T = 2$ . Let *X* denote the cost of replacement parts. It is assumed to be gamma distributed with parameters  $\alpha$  and  $\theta_X$  such that:

$$100 = E[X] = \alpha \theta_X$$

$$5,000 = \operatorname{var}(X) = \alpha \theta_X^2$$

Solving these equations results in:

$$\alpha = 2$$
  $\theta_X = 50$ 

We are asked to calculate the probability of the event  $\{T > 3\} \cup \{X > 150\}$ .

We can calculate the two components as follows:

$$\Pr(T > 3) = 1 - \Pr(T < 3) = e^{-3/2} = 0.2231$$

$$\Pr(X > 150) = 1 - \Pr(X < 150) = e^{-150/50} \left(1 + \frac{150}{50}\right) = 0.1991$$

By the additive probability law, we have:

$$Pr(\{T>3\} \cup \{X>150\}) = Pr(T>3) + Pr(X>150) - Pr(\{T>3\} \cap \{X>150\})$$
$$= Pr(T>3) + Pr(X>150) - Pr(T>3)Pr(X>150)$$
$$= 0.2231 + 0.1991 - 0.2231 \times 0.1991 = 0.3778$$

#### Solution 6.10

We can calculate the required probability as follows:

$$\Pr(4 < S < 8) = \Pr(4 < S < 8 \cap N = 0) + \Pr(4 < S < 8 \cap N = 1) + \Pr(4 < S < 8 \cap N > 1)$$

If there are no claims (ie N = 0), then the claim amount must be zero, so:

$$\Pr(4 < S < 8 \cap N = 0) = 0$$

Therefore:

$$\Pr(4 < S < 8) = \Pr(4 < S < 8 \cap N = 1) + \Pr(4 < S < 8 \cap N > 1)$$

So from basic laws of probability, we have:

$$Pr(4 < S < 8) = Pr(4 < S < 8 \cap N = 1) + Pr(4 < S < 8 \cap N > 1)$$
  
=  $Pr(4 < S < 8 | N = 1) \times Pr(N = 1) + Pr(4 < S < 8 | N > 1) \times Pr(N > 1)$   
=  $\underbrace{(F(8) - F(4))}_{exp, \theta=5} \times \frac{1}{3} + \underbrace{(F(8) - F(4))}_{exp, \theta=8} \times \frac{1}{6}$   
=  $\left((1 - e^{-8/5}) - (1 - e^{-4/5})\right) \times \frac{1}{3} + \left((1 - e^{-8/8}) - (1 - e^{-4/8})\right) \times \frac{1}{6}$   
= 0.1223

The waiting time between accidents follows an exponential distribution with mean  $\theta = \frac{5}{2} = 2.5$  days.

Hence:

$$\Pr(X > 3) = 1 - F(3) = e^{-\frac{3}{2.5}} = e^{-1.2} = 0.3012$$

#### Solution 6.12

From the form of the pdf, we can see that *X* follows a gamma distribution with  $\alpha = 6$  and  $\theta = 100$ . Hence:

$$E[X] = \alpha \theta = 6 \times 100 = 600$$

#### Solution 6.13

First, we'll determine the parameter values:

$$15 = E[X] = \alpha \theta$$
  

$$75 = \operatorname{var}(X) = \alpha \theta^{2}$$
  

$$\Rightarrow \theta = \frac{\operatorname{var}(X)}{E[X]} = 5 \Rightarrow \alpha = 3$$

Since  $\alpha$  is a positive integer, the cdf can be written as:

$$F(x) = 1 - e^{-x/5} \left( 1 + \frac{x}{5} + \frac{x^2}{2!5^2} \right)$$

Finally:

$$\Pr(X > 30 | X > 15) = \frac{\Pr(X > 30)}{\Pr(X > 15)} = \frac{1 - F(30)}{1 - F(15)} = \frac{1 - 0.93803}{1 - 0.57681} = \frac{0.06197}{0.42319} = 0.1464$$

#### Solution 6.14

From the form of the moment generating function, we can see that *X* follows a gamma distribution with parameters  $\alpha = 2$  and  $\theta = 3$ . Hence:

$$\operatorname{var}(X) = \alpha \theta^2 = 2 \times 3^2 = 18$$

Differentiating the cumulant generating function, we have:

$$R_X(t) = \ln M_X(t) = \ln \left( (1 - \theta t)^{-\alpha} \right) = -\alpha \ln (1 - \theta t)$$

$$R'_X(t) = \frac{\alpha \theta}{(1 - \theta t)}$$

$$R''_X(t) = \frac{\alpha \theta^2}{(1 - \theta t)^2} \implies R''_X(0) = \alpha \theta^2$$

$$R'''_X(t) = \frac{2\alpha \theta^3}{(1 - \theta t)^3} \implies R'''_X(0) = 2\alpha \theta^3$$

Hence the skewness is:

$$\frac{E[(X-\mu)^3]}{\sigma^3} = \frac{R_X''(0)}{(R_X''(0))^{3/2}} = \frac{2\alpha\theta^3}{\alpha^{3/2}\theta^3} = \frac{2}{\sqrt{\alpha}}$$

#### Solution 6.16

 $\Rightarrow$  From 2010, this question is no longer covered by the Exam P syllabus.

#### Solution 6.17

 $\Rightarrow$  From 2010, this question is no longer covered by the Exam P syllabus.

#### Solution 6.18

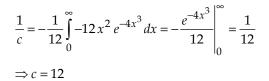
First we need the value of c. This can be found from:

$$1 = \int_{0}^{\infty} c x^2 e^{-4x^3} dx$$
$$\Rightarrow \frac{1}{c} = \int_{0}^{\infty} x^2 e^{-4x^3} dx$$

Now:

$$\int g'(x) e^{g(x)} dx = e^{g(x)} + k$$

where k is a constant.



Now we need:

$$\Pr(X > 0.5) = \int_{0.5}^{\infty} 12 x^2 e^{-4x^3} dx = -e^{-4x^3} \Big|_{0.5}^{\infty} = e^{-4(0.5)^3} = e^{-0.5} = 0.6065$$

#### Solution 6.19

 $\Rightarrow$  From 2010, this question is no longer covered by the Exam P syllabus.

#### Solution 6.20

 $\Rightarrow$  From 2010, this question is no longer covered by the Exam P syllabus.